Perfect simulation and coupling from the past

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Markov chain

- Consider a finite number of *states*: $S = \{s_1, s_2, s_3\}$.
- Next, let a *transition matrix* **P** describe how likely it is to move between the different states:



| | <i>p</i> ₁₁ | p ₁₂ | <i>p</i> ₁₃ |
|------------|------------------------|------------------------|------------------------|
| P = | <i>p</i> ₂₁ | <i>p</i> ₂₂ | <i>p</i> ₂₃ |
| | _p ₃₁ | <i>p</i> ₃₂ | <i>p</i> ₃₃ |

- A Markov chain is memoryless
 - Given the present, the future does not depend on the past.

Markov chain, equilibrium distribution

For a Markov chain $\{X_1, X_2, \dots, X_n\}$ with transition matrix **P**:

•
$$X_t \sim \pi^T \Rightarrow X_{t+1} \sim \pi^T P$$
.

• $\lim_{n\to\infty} \pi_{(1)}^T \mathbf{P}^n \to \pi^*$ - π^* is the equilibrium distribution of the Markov chain.



Monte Carlo

- Suppose you are able to simulate $X_1, X_2, ..., X_n$, which are all independent and identically distributed (i.i.d.).
- You wish to calculate the expectation $\mu = \mathsf{E}[X]$
- If exact calculation is infeasible, we can use a Monte Carlo estimate:

$$\hat{\mu}_n = \frac{1}{n} \sum_{i=1}^n X_i$$

- The expectation is approximated by the empirical mean of independent samples (law of large numbers)
- The statistical accuracy is inversely proportional to the square root of the sample size.
 - Increasing the sample size by a factor of 100 reduces the estimation error by a factor of 10.

Markov chain Monte Carlo (MCMC)

- Not able to generate i.i.d. samples from the desired distribution.
 - Construct a Markov chain that has the target distribution as its equilibrium distribution.
- Simulate (correlated) state values $\tilde{X_1}, \tilde{X_2}, \dots, \tilde{X_n}$ from the Markov chain.
- Apply standard Monte Carlo to the Markov chain output:

$$\hat{\mu}_n = \frac{1}{n} \sum_{i=1}^n \tilde{X}_i$$

MCMC example, discrete distribution

- We use the Markov chain to model the weather: $S = \{s_1 = rainy, s_2 = sunny, s_3 = cloudy\}.$
- $P(rain) \approx$ The proportion of simulated time periods spent in s_1 .
- As we have seen, this simple Markov chain quickly reaches its equilibrium distribution. However, Monte Carlo error is still present.



Markov chain, continuous distribution

• Three different runs of the same Markov chain, started with different initial values.



When has the Markov chain reached equilibrium?

- The chain must be run long enough so that the result is independent of the starting state.
 - May be possible to derive an upper bound of iterations needed.
 - Employ various convergence diagnostics on the simulation output.
- A popular diagnostic is to compare the within-chain variances to the variance of all the chains mixed together.
 - At convergence, these variances should be (close to) identical.
- Diagnostics can identify lack of convergence, but they do not prove convergence.

- In 1996, Propp and Wilson showed how MCMC algorithms can be modified so that they deliver *exact* draws from the Markov chains' equilibrium distribution.
- This type of sampling is referred to as perfect simulation.
- Perfect sampling determines automatically how long the Markov chain must run to reach the **exact** equilibrium distribution.
 - Ensures unbiased estimates.
 - Must still account for Monte Carlo error.

- The algorithm of Propp and Wilson is called Coupling from the past.
- Their simple, yet revolutionary, idea is to run the Markov chain **from the past** instead of into the future.
- Into the future: X_0 , X_1 , X_2 , \ldots , X_{n-2} , X_{n-1} , X_n
- From the past: $X_{-n}, X_{-n+1}, X_{-n+2}, \ldots, X_2$, X_1 , X_0
- At first glance, this is simply a re-labelling of the Markov chain.
 How is it helpful?

• Instead of describing a Markov chain with a transition matrix, we can use an *update function*.

 $- X_{t+1} = \phi(X_t, \varepsilon_{t+1})$

 Letting the distribution of all ε_t be independent standard uniform, our three-state example get the following update function:.

$$\boldsymbol{P} = \begin{bmatrix} p_{11} & p_{12} & p_{13} \\ p_{21} & p_{22} & p_{23} \\ p_{31} & p_{32} & p_{33} \end{bmatrix} \quad \phi(s_i, \varepsilon) = \begin{cases} 1, \text{ for } \varepsilon \in [0, p_{i1}) \\ 2, \text{ for } \varepsilon \in [p_{i1}, p_{i1} + p_{i2}) \\ 3, \text{ for } \varepsilon \in [p_{i1} + p_{i2}, 1]. \end{cases}$$

• Unlike the transition matrix, the update function explicitly describes how the next state depends on the simulated random variable.

Into the future vs. from the past

- Into the future: X_0 , X_1 , X_2 , \ldots , X_{n-2} , X_{n-1} , X_n
- From the past: $X_{-n}, X_{-n+1}, X_{-n+2}, ..., X_{-2}, X_{-1}, X_0$
- Into the future: $(X_n|X_0 = x) = \phi(\phi(\dots,\phi(\phi(x,\varepsilon_1),\varepsilon_2),\dots,\varepsilon_{n-1}),\varepsilon_n)$
- From the past: $(X_0|X_{-n} = x) = \phi(\phi(\dots,\phi(\phi(x,\varepsilon_{-n+1}),\varepsilon_{-n+2}),\dots,\varepsilon_{-1}),\varepsilon_0)$
- The sequences have identical distribution, given that ε_t is i.i.d., which typically is the case.
- This means that the limits as $n \to \infty$ also have the same distribution:

$$- \lim_{n\to\infty} (X_n|X_0=x) \sim \lim_{n\to\infty} (X_0|X_{-n}=x)$$

- The limit is impossible to compute when simulating into the future. Why should it be easier to calculate when simulating from the past?
- A simple example shows why. Assume $\phi(X_t, \varepsilon_{t+1}) = \varepsilon_{t+1}$.
 - With the standard i.i.d assumption for ε_t , t = 1, 2, ..., it is clear this Markov chain forms an i.i.d. sequence.
- Into the future:
 - $-X_0, \varepsilon_1, \varepsilon_2, \dots, \varepsilon_n$ $-\lim_{n \to \infty} (X_n | X_0 = x) \sim \epsilon_0$
- From the past:
 - $\begin{array}{l} X_{-n}, \varepsilon_{-n+1}, \varepsilon_{-n+2}, \dots, \varepsilon_{0} \\ \lim_{n \to \infty} (X_{0} | X_{-n} = x) = \epsilon_{0} \end{array}$

- Given certain conditions, it can be shown that there exists a finite stopping time T such that the distribution of $(X_0|X_{-T} = x)$ is exactly the equilibrium distribution of the Markov chain.
- The algorithm start by defining a sequence of "starting times".
 Typically {-1, -2, -4, -8, -16, ... }.
- The goal of the algorithm is to find a starting time -T, where multiple chains with different initial states end up in the same state at t = 0.

Coupling from the past (CFTP)

- $\phi(s_1, \varepsilon_0) = s_1$
- $\phi(s_2, \varepsilon_0) = s_2$
- $\phi(s_3, \varepsilon_0) = s_1$





Coupling from the past (CFTP)

- $\phi(s_1, \varepsilon_{-2}) = s_2$
- $\phi(s_2, \varepsilon_{-2}) = s_1$
- $\phi(s_3, \varepsilon_{-2}) = s_3$

φ(s₁, ε₋₃) = s₂
φ(s₂, ε₋₃) = s₃

•
$$\phi(s_3, \varepsilon_{-3}) = s_3$$



CFTP extension - Sandwiching

• For the CFTP algorithm to be of practical use, it needs to work in cases where the state-space S of the Markov chain is very large.



• Read-once CFTP (Wilson)

- Requires less computer memory, as it does not need to store all the random variables ε_t .

Fill's algorithm

- Makes it possible to abandon chains that runs too long, without introducing sampling bias.
- Small-set CFTP (Green & Murdoch)
 - Can be applied to continuous state-spaces.

- Perfect sampling determines automatically how long the Markov chain must run to reach the **exact** equilibrium distribution.
- The most successful applications of perfect sampling involves large, discrete state-spaces.
- Unlike standard MCMC, there is no universal method for perfect sampling that is suitable for all Markov chains.

Bibliography



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